

MATH 829: Introduction to Data Mining and Analysis

The EM algorithm

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Some strategies for dealing with missing values

Some options for dealing with missing values:

- **Deletion** (delete observations, remove variable, etc.). Solves the problem, but ignores some of the data (can be significant). May lead to ignoring an entire "category" of observations. Can generate significant bias.
- **Interpolation**. Sometimes it is possible to interpolate missing values (e.g. timeseries). However, we need enough data to be able to produce a good interpolation. In some problems, interpolation is not an option (e.g. age in the titanic passenger data).
- **Replace missing value with mean**. May introduce bias. Only valid for numerical observations.
- **Imputation with the EM algorithm**. Replace missing values by the *most likely values*. Account for all information available. Much more rigorous. However, requires a model. Can be computationally intensive.

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Missing values in data

Missing data is a common problem in statistics.

- No measurement for a given individual/time/location, etc.
- Device failed.
- Error in data entry.
- Data was not disclosed for privacy reasons.
- etc.

```
Swandercock, Mr. William Henry    male  20.0  0
Andersson, Mr. Anders Johan      male  39.0  1
Vestrom, Miss. Hulda Amanda Adolfina female  14.0  0
Howlett, Mrs. (Mary D Kingcome)  female  55.0  0
Rice, Master, Eugene             male   7.0  4
Williams, Mr. Charles Eugene     male   NaN  0
Vander Planke, Mrs. Julius (Emelia Maria Vande... female  31.0  1
Musselwhite, Mrs. Fatima         female  NaN  0
```

Source: www.kaggle.com/mikekaggle/titanic.

How can we deal with missing values?

- Many possible procedures.
- The choice of the procedure can significantly impact the conclusions of a study.

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Missing data mechanism

"Types" of missing data:

- **Missing completely at random (MCAR)**: The events that lead to a missing value are independent both of *observable variables* and of the *unobservable parameters* of interest, and occur entirely at random. (Rarely the case in practice.)
- **Missing at random (MAR)**: missingness is not random, but can be fully accounted for by *observed values*.
- **Missing not at random (MNAR)**: neither MAR nor MCAR.

Example: a study about people's weight. We measure (weight, sex).

- Some respondent may not answer the survey for no particular reason. MCAR
- Maybe women are less likely to answer than male (independently of their weight). MAR
- Heavy or light people may be less likely to disclose their weight. MNAR.

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Example

- Suppose we have **independent** observations of a *discrete* random vector $X = (X_1, X_2, X_3, X_4)$ taking values in $\{0, 1, 2, 3\}$.

X_1	X_2	X_3	X_4
2	0	2	3
3	NA	1	1
1	3	NA	NA
2	NA	1	NA

- Let $p(x_1, x_2, x_3, x_4) = P(X_1 = x_1, X_2 = x_2, X_3 = x_3, X_4 = x_4)$ be the pmf of X .
- Ignoring the missing data mechanism, we have

$$p(x_1, \text{NA}, x_3, x_4) = \sum_{x_2=0}^3 p(x_1, x_2, x_3, x_4).$$

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Example (cont.)

- Suppose the data comes from a parametric model $p(x_1, x_2, x_3, x_4; \theta)$ where $\theta \in \Theta$ is unknown.

X_1	X_2	X_3	X_4
2	0	2	3
3	NA	1	1
1	3	NA	NA
2	NA	1	NA

- We compute the *likelihood* of the data:

$$L(\theta) = p(2, 0, 2, 3) \times p_{1,3,4}(3, 1, 1) \times p_{1,2}(1, 3) \times p_{1,3}(2, 1),$$

where $p_{1,3,4}(x_1, x_3, x_4) = \sum_{x_2=0}^3 p(x_1, x_2, x_3, x_4)$, $p_{1,2}(x_1, x_2) = \sum_{x_3=0}^3 \sum_{x_4=0}^3 p(x_1, x_2, x_3, x_4)$, and $p_{1,3}(x_1, x_3) = \sum_{x_2=0}^3 \sum_{x_4=0}^3 p(x_1, x_2, x_3, x_4)$ denote *marginals* of p .

- The *likelihood* can now be maximized as a function of θ .

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Imputing the missing values

- Recall that $f(x) = E(Y|X = x)$ has the following optimality property:

$$E(Y|X = x) = \underset{c \in \mathbb{R}}{\operatorname{argmin}} E(Y - c)^2$$

where c is some function of x .

- So $E(Y|X = x)$ is the "best prediction" of Y given X in the mean squared error sense.
- As a result, once $p(x; \theta)$ is known (after estimating θ by maximum likelihood for example), we can *impute* missing values using:

$$\hat{x}_{\text{miss}} = E(x_{\text{miss}} | x_{\text{observed}}).$$

For example, if $x = (1, 3, \text{NA}, \text{NA})$ then:

$$(\hat{x}_3, \hat{x}_4) = E((X_3, X_4) | X_1 = 1, X_2 = 3),$$

where E is computed with respect to $p(x_1, x_2, x_3, x_4; \theta)$.

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Summary

In summary, given a family of probability models $p(x; \theta)$ for the data, under MAR, we can:

- Compute the likelihood of θ by *marginalizing* over the missing values.
- Estimate the parameter θ by maximum likelihood.
- Impute missing values using $\hat{x}_{\text{miss}} = E_{\theta}(x_{\text{miss}} | x_{\text{observed}})$, where E_{θ} denotes the expected value with respect to the probability distribution p_{θ} .

Remark: We assumed above that the variables are discrete, and the observations are independent for simplicity. The same procedure applied in the general case.

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- The methodology described so far solves our missing data problem in principle.
- However, explicitly finding the maximum of the likelihood function can be very difficult.

The **Expectation-Maximization (EM) algorithm** of *Dempster, Laird, and Rubin, 1977* provides a more efficient way of solving the problem.

The EM algorithm leverages the fact the the likelihood is often easy to maximize if there is no missing values.

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For simplicity, we will assume our observations are independent and the random variables are discrete.

Some notation:

- We have a random vector W taking values in \mathbb{R}^P .
- The distribution of the vector is $p(w; \theta)$.
- We want to estimate θ .
- We only observe a part of the vector

$$(x^{(i)}, z^{(i)}) \in \mathbb{R}^{P_x} \times \mathbb{R}^{P-z} \quad (i = 1, \dots, n).$$

- So $x^{(i)}$ is the **observed** part and $z^{(i)}$ is the **unobserved** part.
- The log-likelihood function is given by

$$l(\theta) = \sum_{i=1}^n \log p(x^{(i)}; \theta) = \sum_{i=1}^n \log \sum_{z^{(i)}} p(x^{(i)}, z^{(i)}; \theta).$$

(the second sum is over all the possible values of $z^{(i)}$).

- We would like to maximize that function over θ (generally difficult).

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The EM algorithm (cont.)

Instead of trying to maximize the log-likelihood directly, the EM algorithm constructs a sequence of approximations $\theta^{(i)}$ of θ .

- Let $\theta^{(0)}$ be an **initial guess** for θ .
- Given the current estimate $\theta^{(i)}$ of θ , compute

$$Q(\theta|\theta^{(i)}) := E_{z|x, \theta^{(i)}} \log p(x, z; \theta) \\ = \sum_{i=1}^n E_{z^{(i)}|x^{(i)}, \theta^{(i)}} \left(\log p(x^{(i)}, z^{(i)}; \theta) \right) \quad (\text{E step})$$

(In other words, we average the missing values according to their distribution after observing the observed values.)

- We then optimize $Q(\theta|\theta^{(i)})$ with respect to θ :

$$\theta^{(i+1)} := \underset{\theta}{\operatorname{argmax}} Q(\theta|\theta^{(i)}) \quad (\text{M step}).$$

Theorem: The sequence $\theta^{(i)}$ constructed by the EM algorithm satisfies:

$$l(\theta^{(i+1)}) \geq l(\theta^{(i)}).$$

Remark: There is no guarantee that the EM algorithm will find the **global** max of the likelihood. It may only find a **local** max.

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