

MATH 829: Introduction to Data Mining and Analysis
Linear Regression: statistical tests

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February 17, 2016

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Statistical hypothesis testing

Suppose we have a linear model for some data:

$$Y = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \epsilon.$$

- An important problem is to identify which variables are really useful in predicting Y .
- We want to decide if $\beta_i = 0$ or not with some level of confidence.
- Also want to test if groups of coefficients $\{\beta_k : k = 1, \dots, l\}$ are zero.

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Statistical hypothesis testing (cont.)

Recall: to do a statistical test:

1. State a null hypothesis H_0 and an alternative hypothesis H_1 .
2. Construct an appropriate test statistics.
3. Derive the distribution of the test statistic under the null hypothesis.
4. Select a significance level α (typically 5% or 1%).
5. Compute the test statistics and decide if the null hypothesis is rejected at the given significance level.

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Example

Example: Suppose $X \sim N(\mu, 1)$ with μ unknown. We want to test:

$$H_0 : \mu = 0$$

$$H_1 : \mu \neq 0.$$

We have an iid sample $X_1, \dots, X_n \sim N(\mu, 1)$.
Recall: if $X \sim N(\mu_1, \sigma_1^2)$ and $Y \sim N(\mu_2, \sigma_2^2)$ are independent, then $X + Y \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$.

Therefore, under H_0 , we have

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n X_i \sim N\left(0, \frac{1}{n}\right).$$

Test statistics: $\sqrt{n} \cdot \hat{\mu} \sim N(0, 1)$.

Suppose we observed: $\sqrt{n}\hat{\mu} = k$.

We compute:

$$P(-z_\alpha \leq \sqrt{n}\hat{\mu} \leq z_\alpha) = P(-z_\alpha \leq N(0, 1) \leq z_\alpha) = 1 - \alpha.$$

Reject the null hypothesis if $k \notin [-z_\alpha, z_\alpha]$.

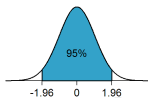
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Example (cont.)

For example, suppose: $\alpha = 0.05$, $\sqrt{n}\hat{\mu} = 2.2$.

If $Z \sim N(0, 1)$, then

$$P(-1.96 \leq Z \leq 1.96) = 0.95$$



Therefore, it is very unlikely to observe $\sqrt{n}\hat{\mu} = 2.2$ if $\mu = 0$. We reject the null hypothesis.

In fact, $P(-2.2 \leq Z \leq 2.2) \approx 0.972$. So our **p-value** is 0.028.

• Type I error: H_0 true, but rejected \rightarrow False positive. (Controlled by the level α).

• Type II error: H_0 false, but not rejected \rightarrow False negative. (Power of the test).

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Testing if coefficients are zero

- In practice, we often include unnecessary variables in linear models.
- These variables bias our estimator, and can lead to poor performance.
- Need ways of identifying a "good" set of predictors.

We now discuss a classical approach that uses statistical tests.

Before, we tested if the mean of a $N(\mu, 1)$ is zero:

$$H_0: \mu = 0$$

$$H_1: \mu \neq 0.$$

assuming $\sigma^2 = 1$ is known. What if the variance is *unknown*?

Sample variance:

$$s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{X})^2, \quad \bar{X} = \frac{1}{n} \sum_{i=1}^n x_i.$$

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Testing if coefficients are zero (cont.)

In general, suppose $X \sim N(\mu, \sigma^2)$ with σ^2 known and we want to test

$$H_0: \mu = \mu_0$$

$$H_1: \mu \neq \mu_0.$$

Under the H_0 hypothesis, we have

$$\bar{X} := \frac{1}{n} \sum_{i=1}^n X_i \sim N\left(\mu_0, \frac{\sigma^2}{n}\right)$$

Therefore, we use the test statistic

$$Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} \sim N(0, 1).$$

If the variance is unknown, we replace σ by its sample version s

$$T = \frac{\bar{X} - \mu_0}{s/\sqrt{n}} \sim t_{n-1}.$$

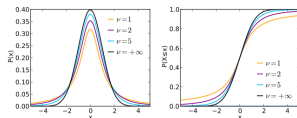
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Review: the student distribution

The student t_ν distribution with ν degrees of freedom:

$$f_\nu(t) = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\sqrt{\nu\pi} \Gamma\left(\frac{\nu}{2}\right)} \left(1 + \frac{t^2}{\nu}\right)^{-\frac{\nu+1}{2}},$$

where Γ is the Gamma function.



When X_1, \dots, X_n are iid $N(\mu, \sigma^2)$, then

$$\frac{\bar{X} - \mu}{s/\sqrt{n}} \sim t_{n-1}.$$

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Testing if coefficients are zero (cont.)

Back to testing regression coefficients: suppose

$$y_i = x_{i,1}\beta_1 + x_{i,2}\beta_2 + \cdots + x_{i,p}\beta_p + \epsilon_i,$$

where (x_{ij}) is a fixed matrix, and ϵ_i are iid $N(0, \sigma^2)$.

We saw that this implies

$$\hat{\beta} \sim N(\beta, (X^T X)^{-1} \sigma^2).$$

In particular,

$$\hat{\beta}_i \sim N(\beta_i, v_i \sigma^2),$$

where v_i is the i -th diagonal element of $(X^T X)^{-1}$.

We want to test:

$$H_0: \beta_i = 0$$

$$H_1: \beta_i \neq 0.$$

Note: v_i is known, but σ is unknown.

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Testing if coefficients are zero (cont.)

Recall:

$$y_i = x_{i,1}\beta_1 + x_{i,2}\beta_2 + \cdots + x_{i,p}\beta_p + \epsilon_i,$$

Problem: How do we estimate σ ?

Let $\hat{\epsilon}_i = y_i - (x_{i,1}\hat{\beta}_1 + x_{i,2}\hat{\beta}_2 + \cdots + x_{i,p}\hat{\beta}_p)$. What about:

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n \hat{\epsilon}_i^2 \quad ?$$

What is $E(\hat{\sigma}^2)$?

We have $y = X\beta + \epsilon$ and $\hat{\beta} = (X^T X)^{-1} X^T y$. Thus,

$$\begin{aligned} \hat{\epsilon} &= y - X\hat{\beta} \\ &= y - X(X^T X)^{-1} X^T y \\ &= (I - X(X^T X)^{-1} X^T) y \\ &= (I - X(X^T X)^{-1} X^T) (X\beta + \epsilon) \\ &= (I - X(X^T X)^{-1} X^T) \epsilon \\ &= M\epsilon. \end{aligned}$$

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Testing if coefficients are zero (cont.)

We showed $\hat{\epsilon} = M\epsilon$ where $M := I - X(X^T X)^{-1} X^T$. Now

$$\sum_{i=1}^n \hat{\epsilon}_i^2 = \hat{\epsilon}^T \hat{\epsilon} = \epsilon^T M^T M \epsilon.$$

Note: $M^T = M$ and

$$\begin{aligned} M^T M &= M^2 = (I - X(X^T X)^{-1} X^T) (I - X(X^T X)^{-1} X^T) \\ &= I - X(X^T X)^{-1} X^T = M. \end{aligned}$$

(In other words, M is idempotent.)

Therefore,

$$\sum_{i=1}^n \hat{\epsilon}_i^2 = \epsilon^T M \epsilon.$$

Now,

$$\begin{aligned} E(\epsilon^T M \epsilon) &= E(\text{tr}(M \epsilon \epsilon^T)) \\ &= \text{tr} E(M \epsilon \epsilon^T) \\ &= \text{tr} M E(\epsilon \epsilon^T) \\ &= \text{tr} M \sigma^2 I = \sigma^2 \text{tr} M. \end{aligned}$$

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Testing if coefficients are zero (cont.)

We proved:

$$E\left(\sum_{i=1}^n \hat{\epsilon}_i^2\right) = \sigma^2 \text{tr} M,$$

where $M = I - X(X^T X)^{-1} X^T$. (Here $I = I_n$, the $n \times n$ identity matrix.)

What is $\text{tr} M$? Recall $\text{tr}(AB) = \text{tr}(BA)$. Thus,

$$\begin{aligned} \text{tr} M &= \text{tr}(I - X(X^T X)^{-1} X^T) \\ &= n - \text{tr}(X(X^T X)^{-1} X^T) \\ &= n - \text{tr}(X^T X (X^T X)^{-1}) \\ &= n - \text{tr}(I_p) \\ &= n - p. \end{aligned}$$

Therefore,

$$\frac{1}{n-p} E\left(\sum_{i=1}^n \hat{\epsilon}_i^2\right) = \sigma^2.$$

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