# MATH 829: Introduction to Data Mining and Analysis <br> Consistency of Linear Regression 

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## Distribution of regression coefficients

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What is the distribution of $\hat{\beta}$ ?

## Multivariate normal distribution

Recall: $X=\left(X_{1}, \ldots, X_{p}\right) \sim N(\mu, \Sigma)$ where

- $\mu \in \mathbb{R}^{p}$,
- $\Sigma=\left(\sigma_{i j}\right) \in \mathbb{R}^{p \times p}$ is positive definite,
if

$$
P(X \in A)=\frac{1}{\sqrt{(2 \pi)^{p} \operatorname{det} \Sigma}} \int_{A} e^{-\frac{1}{2}(x-\mu)^{T} \Sigma^{-1}(x-\mu)} d x_{1} \ldots d x_{p}
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If $Y=c+B X$, where $c \in \mathbb{R}^{p}$ and $B \in \mathbb{R}^{m \times p}$, then

$$
Y \sim N\left(c+B \mu, B \Sigma B^{T}\right)
$$

## Distribution of the regression coefficients (cont.)

Back to our problem: $Y=X \beta+\epsilon$ where $\epsilon_{i}$ are iid $N\left(0, \sigma^{2}\right)$. We have

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Y \sim N\left(X \beta, \sigma^{2} I\right)
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Therefore,

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\hat{\beta}=\left(X^{T} X\right)^{-1} X^{T} Y \sim N\left(\beta, \sigma^{2}\left(X^{T} X\right)^{-1}\right)
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In particular,

$$
E(\hat{\beta})=\beta
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Thus, $\hat{\beta}$ is unbiased.

## Statistical consistency of least squares

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A sequence of estimators $\left\{\theta_{n}\right\}_{n=1}^{\infty}$ of a parameter $\theta$ is said to be consistent if $\theta_{n} \rightarrow \theta$ in probability $\left(\theta_{n} \xrightarrow{p} \theta\right)$ as $n \rightarrow \infty$.


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(Recall: $\theta_{n} \xrightarrow{p} \theta$ if for every $\epsilon>0$,

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In order to prove that $\hat{\beta}_{n}$ (estimator with $n$ samples) is consistent, we will make some assumptions on the data generating model. (Without any assumptions, nothing prevents the observations to be all the same for example...)

## Statistical consistency of least squares (cont.)

Observations: $y=\left(y_{i}\right) \in \mathbb{R}^{n}, X=\left(x_{i j}\right) \in \mathbb{R}^{n \times p}$.

## Statistical consistency of least squares (cont.)

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\begin{aligned}
& \text { Observations: } y=\left(y_{i}\right) \in \mathbb{R}^{n}, X=\left(x_{i j}\right) \in \mathbb{R}^{n \times p} \text {. Let } \\
& \mathbf{x}_{i}:=\left(x_{i, 1}, \ldots, x_{i, n}\right) \in \mathbb{R}^{p} \quad(i=1, \ldots, n) .
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Observations: $y=\left(y_{i}\right) \in \mathbb{R}^{n}, X=\left(x_{i j}\right) \in \mathbb{R}^{n \times p}$. Let
$\mathbf{x}_{i}:=\left(x_{i, 1}, \ldots, x_{i, n}\right) \in \mathbb{R}^{p} \quad(i=1, \ldots, n)$.
We will assume:
(1) $\left(\mathbf{x}_{i}\right)_{i=1}^{n}$ are iid random vectors.
(2) $y_{i}=\beta_{1} x_{i, 1}+\cdots+\beta_{p} x_{i, p}+\epsilon_{i}$ where $\epsilon_{i}$ are iid $N\left(0, \sigma^{2}\right)$.
(3) The error $\epsilon_{i}$ is independent of $\mathbf{x}_{i}$.
(9) $E x_{i j}^{2}<\infty$ (finite second moment).
(3) $Q=E\left(\mathbf{x}_{i} \mathbf{x}_{i}^{T}\right) \in \mathbb{R}^{p \times p}$ is invertible.

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Under these assumptions, we have the following theorem.
Theorem: Let $\hat{\beta}_{n}=\left(X^{T} X\right)^{-1} X^{T} y$. Then, under the above assumptions, we have

$$
\hat{\beta}_{n} \xrightarrow{p} \beta .
$$

## Background for the proof

Recall:
Weak law of large numbers: Let $\left(X_{i}\right)_{i=1}^{\infty}$ be iid random variables with finite first moment $E\left(\left|X_{i}\right|\right)<\infty$. Let $\mu:=E\left(X_{i}\right)$. Then

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\bar{X}_{n}:=\frac{1}{n} \sum_{i=1}^{n} X_{i} \xrightarrow{p} \mu
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Continuous mapping theorem: Let $S, S^{\prime}$ be metric spaces. Suppose $\left(X_{i}\right)_{i=1}^{\infty}$ are $S$-valued random variables such that $X_{i} \xrightarrow{p} X$. Let $g: S \rightarrow S^{\prime}$. Denote by $D_{g}$ the set of points in $S$ where $g$ is discontinuous and suppose $P\left(X \in D_{g}\right)=0$. Then $g\left(X_{n}\right) \xrightarrow{p} g(X)$.

## Proof of the theorem

We have

$$
\hat{\beta}=\left(X^{T} X\right)^{-1} X^{T} y=\left(\frac{1}{n} \sum_{i=1}^{n} \mathbf{x}_{i} \mathbf{x}_{i}^{T}\right)^{-1}\left(\frac{1}{n} \sum_{i=1}^{n} \mathbf{x}_{i} y_{i}\right) .
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Using Cauchy-Schwarz,

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E\left(\left|x_{i j} x_{i k}\right|\right) \leq\left(E\left(x_{i j}^{2}\right) E\left(x_{i k}^{2}\right)\right)^{1 / 2}<\infty .
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By the weak law of large numbers, we obtain

$$
\begin{aligned}
& \frac{1}{n} \sum_{i=1}^{n} \mathbf{x}_{i} \mathbf{x}_{i}^{T} \xrightarrow{p} E\left(\mathbf{x}_{i} \mathbf{x}_{i}^{T}\right)=Q \\
& \frac{1}{n} \sum_{i=1}^{n} \mathbf{x}_{i} y_{i} \xrightarrow{p} E\left(\mathbf{x}_{i} y_{i}\right)
\end{aligned}
$$

Using the continuous mapping theorem, we obtain

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\hat{\beta}_{n} \xrightarrow{p} E\left(\mathbf{x}_{i} \mathbf{x}_{i}^{T}\right)^{-1} E\left(\mathbf{x}_{i} y_{i}\right) .
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(define $g: \mathbb{R}^{p \times p} \times \mathbb{R}^{p} \rightarrow \mathbb{R}^{p}$ by $g(A, b)=A^{-1} b$.)

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Taking expectations,

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We conclude that

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\beta=E\left(\mathbf{x}_{i} \mathbf{x}_{i}^{T}\right)^{-1} E\left(\mathbf{x}_{i} y_{i}\right)
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and so $\hat{\beta}_{n} \xrightarrow{p} \beta$.

